

# The Mass Distribution of Stellar-Mass Black Holes

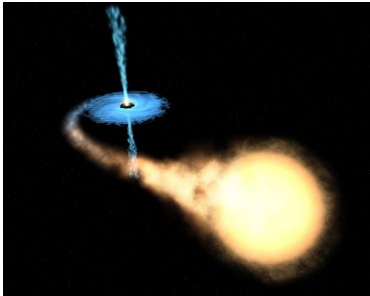
Robust Constraints Through Model Selection  
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# Stellar Mass Black Holes

Black holes come in many varieties; we're interested in the smallest ones. Our sample comes from identified X-Ray Binaries:



## Previous Work

- Bailyn, *et al*, 1998
  - Seven systems.
  - Black hole distribution flat between 5 and 10  $M_{\odot}$ .
  - Mass gap between low-mass BH and high-mass NS.
- Ozel, *et al*, 2010
  - 16 systems.
  - Narrow (Gaussian) distribution with peak at 8  $M_{\odot}$ , width of few.
  - Also mass gap.
  - Also considered theoretically favored models (Freyer and Kalogera 2001).
  - No model comparison.

# Black Hole Masses

We want (accurate) estimates of the mass, so we only use binaries where we know the mass function:

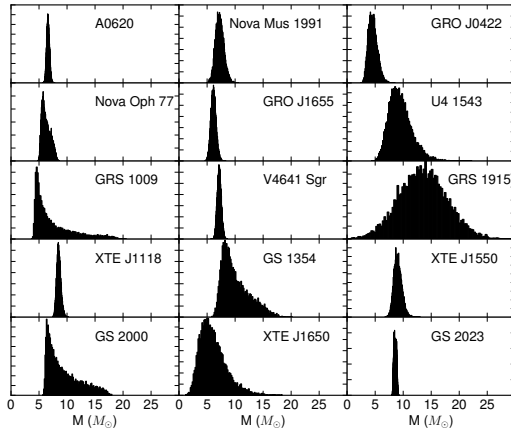
$$f(M) = \frac{PK^3}{2\pi G} = \frac{M \sin^3 i}{(1+q)^2},$$

- $P$  is the period.
- $K$  is the secondary's velocity semi-amplitude.
- $i$  is the inclination.
- $q$  is the mass ratio  $M_2/M$ .

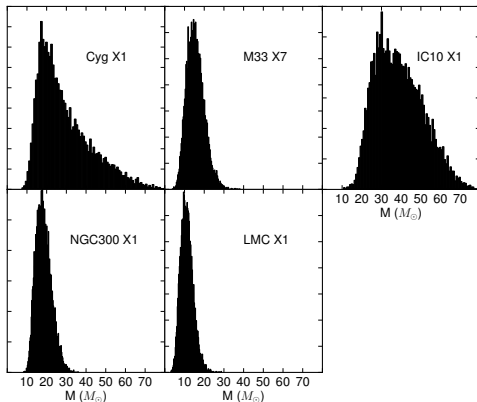
Requires optical observations (spectra) and lightcurve fitting.



# Low Mass Distributions



# High Mass Distributions



# Models for Mass Distribution

Model	$p(M \vec{\theta})$	Params ( $\vec{\theta}$ )
Power Law	$p(M \vec{\theta}) \propto M^\alpha, M_{\min} \leq M \leq M_{\max}$	$\alpha, M_{\min}, M_{\max}$
Gaussian	$p(M \vec{\theta}) = N(\mu, \sigma, M)$	$\mu, \sigma$
Exponential	$p(M \vec{\theta}) \propto \exp\left(-\frac{M}{M_0}\right), M_{\min} < M$	$M_0, M_{\min}$
Log Normal	$p(M \vec{\theta}) \propto \frac{N(\mu, \sigma, \log M)}{M}$	$\mu, \sigma$
Two Gaussian	$p(M \vec{\theta}) \propto \alpha N_1 + (1 - \alpha) N_2$	$\mu_1, \sigma_1,$ $\mu_2, \sigma_2, \alpha$
Histogram	$p(M \vec{\theta}) \propto (\text{bin width})^{-1}$	bin locations



## Models for Mass Distribution II

Each distribution tries to capture properties of  $p(M|d)$ :

**Power Law** Why not? (Can capture  $M_{\min}$  and  $M_{\max}$ .)

**Gaussian** A bump.

**Exponential** Theoretically preferred.

**Log Normal** Long tail.

**Two Gaussian** Secondary bump.

**Histogram** Non-parametric.

# Likelihood

Assume:

- Observations independent.
- Errors on  $f(M)$ ,  $q$ , and  $i$  independent.
- Ignore selection effects

$$p(d|\vec{\theta}) = \prod_{\text{systems}} \int dM p_{\text{obs}}(M) p(M|\vec{\theta})$$

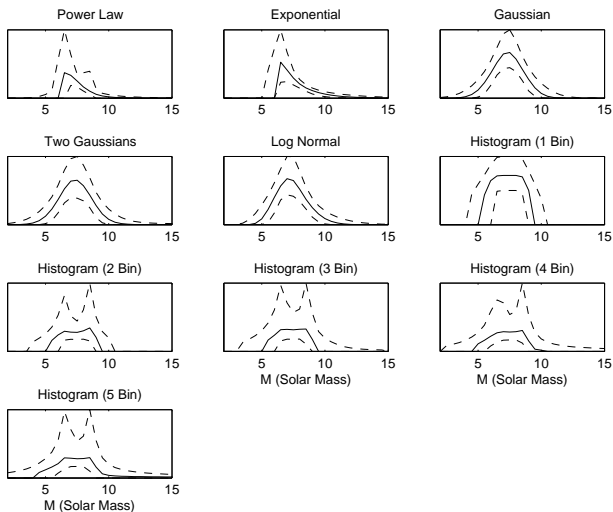
# MCMC

Technique for sampling parameter space proportional to posterior probability:

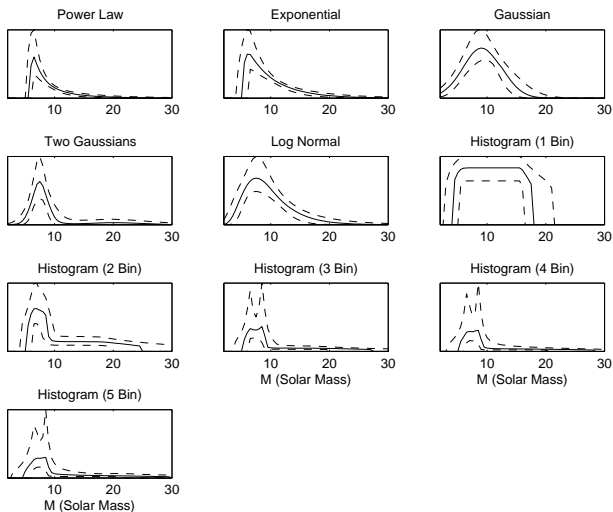
$$n(\vec{\theta}) \propto p(\vec{\theta}|d) = \frac{p(d|\vec{\theta})p(\vec{\theta})}{p(d)}$$

- Wander stochastically in parameter space.
- Propose local jumps, accept/reject probabilistically.
- Obtain a sequence of parameter values:  $\{\vec{\theta}_i | i = 1, \dots\}$
- Each sample represents one particular mass distribution favored by the data.
- Chain is a *distribution of distributions*.

# Low Mass: Distributions of Distributions



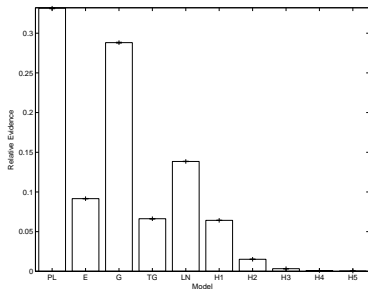
# Complete Sample: Distributions of Distributions



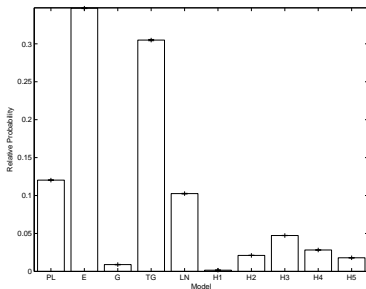
# Model Selection

- Models look similar, but which ones are *the best*?
- Settle with another MCMC, this time jumping *between* models.
- Model counts give relative model probability.

Low-mass:



High-mass:

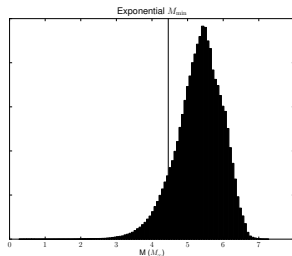
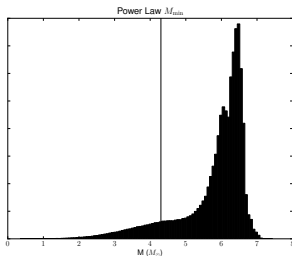


# Minimum BH Mass

Strong evidence for a mass gap in both samples:

Low-mass systems:

Combined sample:

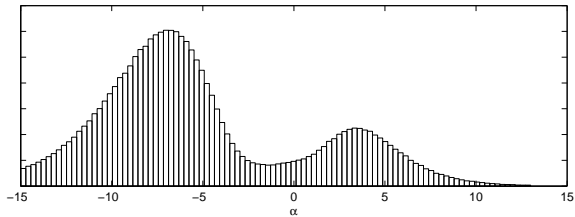
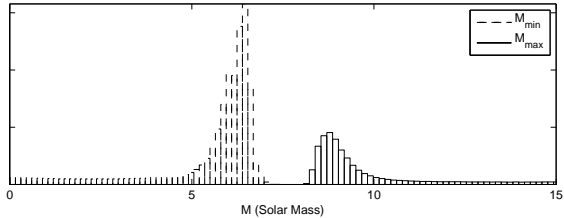


# Conclusion

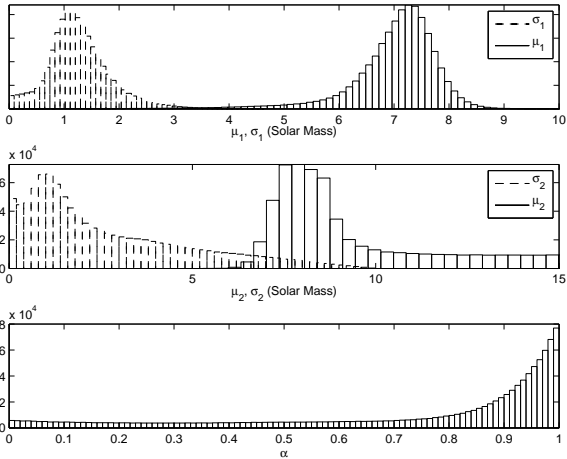
- Best models show large “mass gap”:  
 $M_{\min} \sim 4.25M_{\odot} > 3M_{\odot}$
- Theoretically preferred model disfavored for low-mass systems, OK for complete sample.
- Difference between favored models for low-mass and complete sample implies different underlying distributions.
- Model selection is a powerful technique for extracting information about best-fitting models in the absence of prior knowledge.



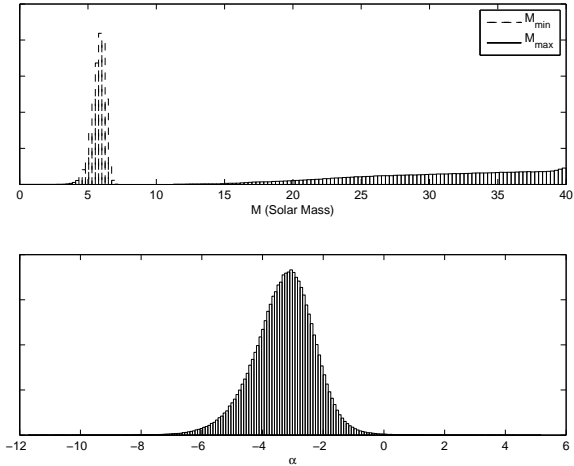
# Power-Law Fitting to a Peak



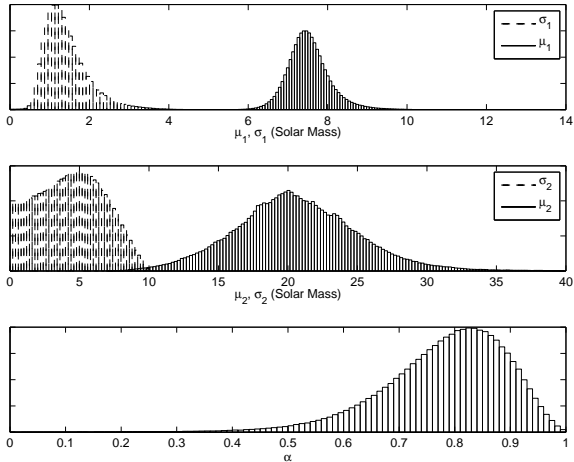
# Two-Gaussian: No Second Bump



# Tail Fixes Power-Law Degeneracy

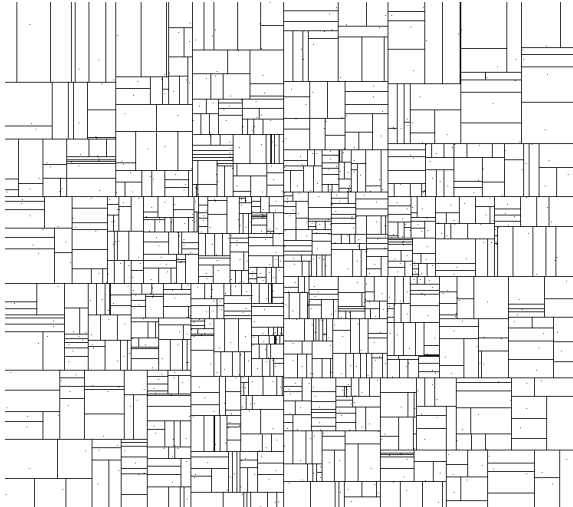


# Two-Gaussian: Second Bump



# Model Selection Techniques

Bayesian discussion on board:



# Model Selection

# Models

We assume a shape for the underlying mass distribution, and then “fit” to the observed masses.... (Bayesian math on board.)

# Power Law

$$p(M|\vec{\theta}) = p(M|\{M_{\min}, M_{\max}, \alpha\}) = \begin{cases} AM^\alpha & M_{\min} \leq m \leq M_{\max} \\ 0 & \text{otherwise} \end{cases}$$



# Decaying Exponential

(Preferred by Fryer & Kalogera (2001))

$$p(M|\vec{\theta}) = p(M|\{M_{\min}, M_0\}) = \begin{cases} \frac{e^{-\frac{M_{\min}}{M_0}}}{M_0} \exp\left[-\frac{M}{M_0}\right] & M_{\min} \leq M \\ 0 & \text{otherwise} \end{cases}$$

# Gaussian and Two Gaussians

$$p(M|\vec{\theta}) = p(M|\{\mu, \sigma\}) = \frac{1}{\sigma\sqrt{2\pi}} \exp \left[ - \left( \frac{M - \mu}{\sqrt{2}\sigma} \right)^2 \right].$$

and

$$p(M|\vec{\theta}) = p(M|\{\mu_1, \mu_2, \sigma_1, \sigma_2, \alpha\}) =$$
$$\frac{\alpha}{\sigma_1\sqrt{2\pi}} \exp \left[ - \left( \frac{M - \mu_1}{\sqrt{2}\sigma_1} \right)^2 \right]$$
$$+ \frac{1 - \alpha}{\sigma_2\sqrt{2\pi}} \exp \left[ - \left( \frac{M - \mu_2}{\sqrt{2}\sigma_2} \right)^2 \right].$$

# Log Normal

$$p(M|\vec{\theta}) = p(M|\{\mu, \sigma\}) = \frac{1}{\sqrt{2\pi}M\sigma} \exp\left[-\frac{(\log M - \mu)^2}{2\sigma^2}\right].$$

# Non-parametric Histograms

Histogram math on board.

# Summary of Models

- Power Law
- Exponential
- Gaussian
- Two Gaussians
- Log Normal
- One-, two-, three-, four-, five-bin histograms